

# Inferring food intake from multiple biomarkers using a latent variable model

Silvia D'Angelo

TCD - 15th March 2023

# Joint work with...

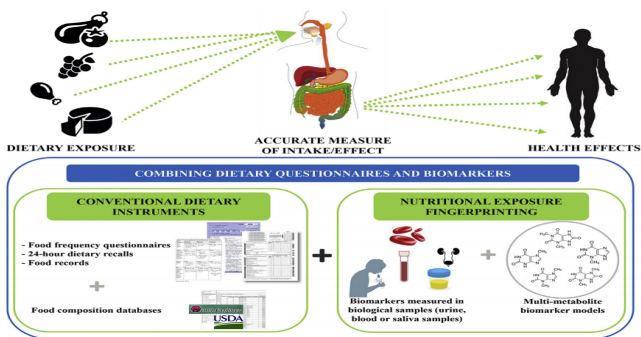


Prof. Claire Gormley



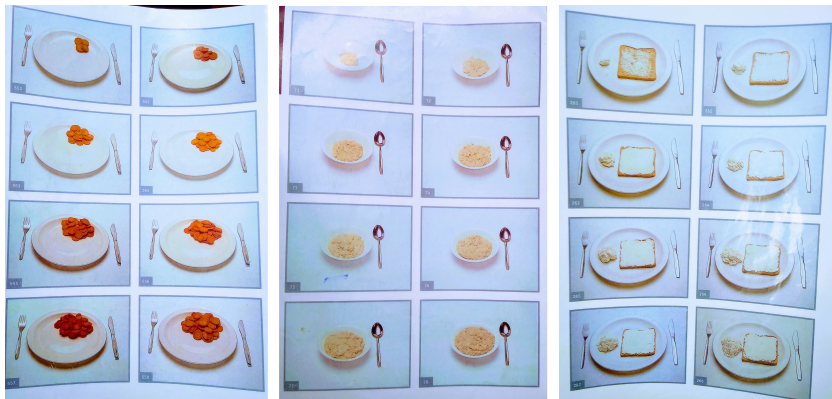
Prof. Lorraine Brennan

# Assessing nutrients and food intake



- Intake data are necessary in various fields (nutrition, epidemiology,...)
- Often the only intake data available are self-reported
- **Self-reported** data are **subjective** and possibly biased
- Dietary **biomarkers** can provide more **objective measurements** of intake

# Portions assessment



# Discovering biomarker candidates

- Experimental studies are designed to investigate the potential of compounds as specific nutrient/food biomarkers
- Participants are fed the food of interest, either directly or indirectly
- Different biological samples are collected and analysed from the participants
- Multiple steps are required to candidate compounds as biomarkers

Experimental studies only investigate the relationship between candidate biomarkers and portions of intake. This should be generalized, to allow biomarkers usage outside of experimental study contexts.

**How to generalize experimental studies results?**

# Previous works on food biomarkers



Gürdeniz et al. 2016.

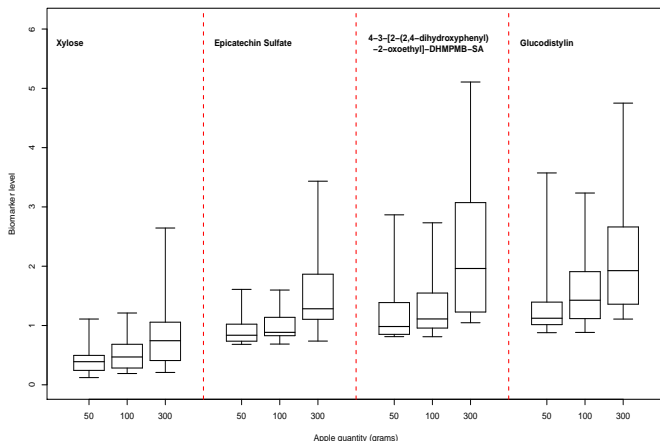
- Consumers
- Non-consumers



Vázquez Manjarrez et al. 2019.

- **Low** intake
- Medium intake
- **High** intake

# Apple intake data



- Feeding study,  $N = 32$  participants
- $D = 3$  apple quantities: 50, 100 and 300 grams
- Collection: 4 days diet + 5<sup>th</sup> day urine + wash-out periods
- Data:  $n = 86$  observations and  $P = 4$  biomarkers

# The model

$$y_{ip} = \alpha_p + \beta_p z_i + \epsilon_{ip},$$

$$z_i \sim \sum_{d=1}^D \pi_{id} \mathcal{N}_{[0,\infty)}(X_d, \theta_d^2), \quad \pi_i = \{\pi_{i1}, \dots, \pi_{iD}\}$$

with

- $P$  biomarkers:  
 $\{y_1, \dots, y_P\}$
- $n$  latent intakes:  
 $\{z_1, \dots, z_n\}$
- $D$  food quantities:  
 $\{X_1, \dots, X_D\}$

and

- $\alpha_p \sim \mathcal{N}_{[0,\infty)}(\mu_\alpha, \sigma_\alpha^2)$
- $\beta_p \sim \mathcal{N}_{(0,\infty)}(\mu_\beta, \sigma_\beta^2)$
- $\epsilon_{ip} \sim \mathcal{N}(0, \sigma_p^2)$
- $y_{ip} \sim \mathcal{N}_{[0,\infty)}(\alpha_p + \beta_p z_i, \sigma_p^2)$



# Modelling the latent intakes

$$z_i \mid c_i \sim \prod_{d=1}^D \left[ \mathcal{N}_{[0, \infty)}(X_d, \theta_d^2) \right]^{[c_i=d]}$$

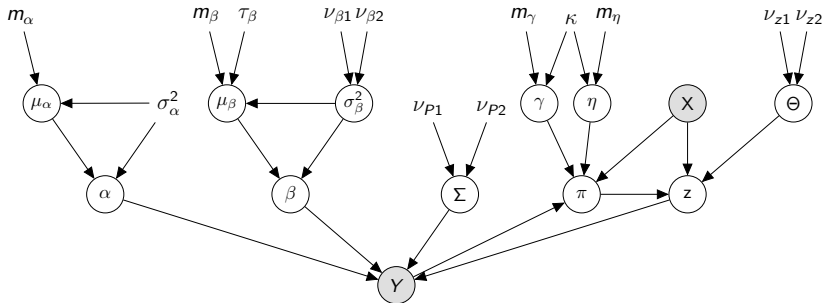
$\mathbf{c} = \{c_1, \dots, c_i, \dots, c_n\}$  are observation-specific allocation labels.

## Modelling the weights

$$\begin{aligned} p(\pi \mid \gamma, \eta, \mathbf{c}, \mathbf{Y}, \mathbf{X}) &= \prod_{i=1}^n \prod_{d=1}^D \pi_{id}^{[c_i=d]} = \prod_{i=1}^n \prod_{d=1}^D \pi_{id}(y_i \mid \gamma_d, \eta)^{[c_i=d]} \\ &= \prod_{i=1}^n \prod_{d=1}^D \left[ \Pr(x_i \leq X_d \mid \gamma_d, \eta, y_i) - \Pr(x_i \leq X_{d-1} \mid \gamma_{d-1}, \eta, y_i) \right]^{[c_i=d]} \end{aligned}$$

where  $\Pr(x_i \leq X_d \mid \gamma_d, \eta, y_i) = \frac{1}{\pi} \left[ \arctan \left( \frac{1}{2}(\gamma_d + \eta y_i) \right) + \frac{\pi}{2} \right]$ .

# Model structure



- $\mu_\alpha \sim \mathcal{N}_{[0,\infty)}(m_\alpha, \tau_\alpha \sigma_\alpha^2)$
- $\mu_\beta \sim \mathcal{N}_{[0,\infty)}(m_\beta, \tau_\beta \sigma_\beta^2)$
- $\sigma_\beta^2 \sim \text{Inv}\Gamma(\nu_{\beta 1}, \nu_{\beta 2})$
- $\theta_d^2 \sim \text{Inv}\Gamma(\nu_{z 1}, \nu_{z 2})$
- $\Sigma = \{\sigma_1^2, \dots, \sigma_p^2\}$
- $\Theta = \{\theta_1^2, \dots, \theta_p^2\}$
- $\gamma_d \sim \mathcal{N}_{(m_{\gamma d-1}, m_{\gamma d+1})}(m_{\gamma d}, \kappa)$
- $\eta_p \sim \mathcal{N}(m_{\eta p}, \kappa)$

# Estimation

- Metropolis within Gibbs MCMC algorithm -

## Outline

- 1 Initialize all parameters
- 2 Gibbs steps: update  $\alpha$ ,  $\beta$ ,  $\mu_\alpha$ ,  $\mu_\beta$ ,  $\sigma_\beta^2$  and  $\Sigma$  from their full conditional distributions
- 3 MH steps: update  $\gamma$  and  $\eta$  (random walk)
- 4 Update labels  $c$
- 5 Gibbs steps: update  $\Theta$  and  $z$  from their full conditional distributions

$$z_i \mid c_i = d, \dots \sim \mathcal{N}_{[0, \infty)}(\mu_{id}^*, \theta_{id}^{2*})$$

$$\theta_{id}^{2*} = \left( \sum_{p=1}^P \frac{\beta_p^2 \theta_d^2 + \sigma_p^2 / P}{\theta_d^2 \sigma_p^2} \right)^{-1}, \mu_{id}^* = \sigma_{id}^{2*} \left[ \sum_{p=1}^P \frac{\beta_p (y_{ip} - \alpha_p)}{\sigma_p^2} + \frac{X_d}{\theta_d^2} \right]$$

# Intake quantification / prediction

**AIM:** Infer intakes for a new group of  $n^*$  observations, when only biomarker measurements are available.

- Sampling distribution for a new intake  $z_j^*$ :

$$p(z_j^* | y_j^*, c_j^*, \Omega) = \prod_{d=1}^D \left[ \mathcal{N}_{[0, \infty)} \left( \frac{\mu_{zj} \theta_d^2 + X_d \sigma_{zj}^2}{\sigma_{zj}^2 + \theta_d^2}, \left( \frac{1}{\theta_d^2} + \frac{1}{\sigma_{zj}^2} \right)^{-1} \right) \right]^{[c_j^* = d]}$$

where:

- $\sigma_{zj}^2 = \left( \sum_{p=1}^P \frac{\beta_p^2}{\sigma_p^2} \right)^{-1}$
- $\mu_{zj} = \sigma_{zj}^2 \left( \sum_{p=1}^P \frac{\beta_p (y_{jp}^* - \alpha_p)}{\sigma_p^2} \right)$
- $\Omega = \{ \alpha, \beta, \Sigma, X, \Theta, \eta, \gamma \}$

- Posterior predictive distribution for a new intake  $z_j^*$ :

$$p(z_j^* | y_j^*, c_j^*) \propto \int p(z_j^* | y_j^*, c_j^*, \Omega) p(\Omega | Y) d\Omega$$

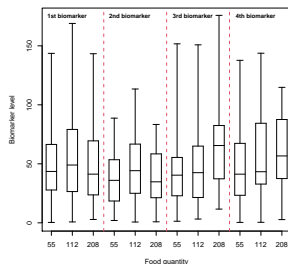
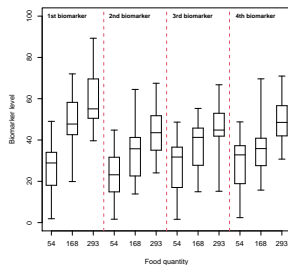
# Simulations - details

## Settings

- $P = 4$  biomarkers
- $n = \{30, 60, 99, 150\}$ ,  $n^* = \lfloor 0.4 \times n \rfloor$
- $(\alpha, \beta)$ : small, medium and large biomarkers range
- $\Sigma$ : small, mixed and large variability
- $X$ : stable, increasing, and decreasing increments
- $\Theta$ : low and high variability
- $c$ : sampled at random in  $\{1, \dots, D\}$

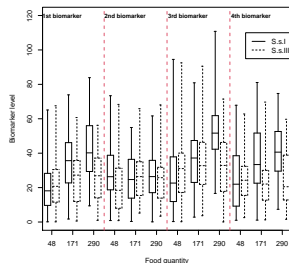
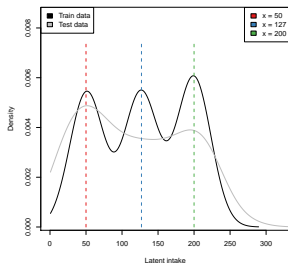
## Outline

- 20 datasets per each settings combination
- 30000 MCMC iterations (burn: 6000)



# Simulations - intake quantification with...

- **Study I:** varying biomarker variability ( $\sigma_p^2$ )
- **Study II:** discrepancies between training and test data generation



- **Study III:** model misspecification

$$y_{ip} = \alpha_p + \beta_p z_i^2 + \epsilon_{ip}$$

# Simulations - results

## Absolute error values (in grams)

between true and estimated/predicted intakes

Sim. study	$\Sigma$	Estimates			Predictions		
		MM	BLR	PLS	MM	BLR	PLS
I	Small	3(7)	77(200)	10(24)	4(9)	76(216)	10(26)
	Mixed	4(8)	62(136)	20(39)	4(33)	111(263)	21(41)
	Large	6(18)	62(136)	35(60)	9(59)	64(137)	37(69)
II	Small	3(8)	61(181)	9(30)	5(25)	62(184)	10(37)
	Mixed	4(8)	113(222)	22(33)	7(22)	112(224)	23(43)
	Large	7(62)	80(118)	39(56)	22(77)	88(108)	41(67)
III	Small	6(35)	66(227)	26(72)	8(49)	67(231)	31(77)
	Mixed	7(44)	98(269)	43(80)	9(56)	112(298)	46(87)
	Large	11(62)	87(197)	64(65)	27(74)	101(138)	70(80)

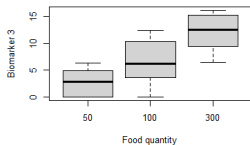
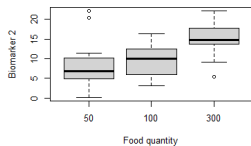
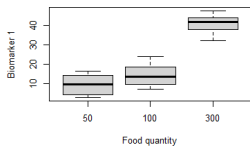
## Models:

- MM:  
multiMarker model
- BLR:  
Bayesian linear regression
- PLS:  
PLS regression

# Simulations - Multinomial weights

$$\pi_{id} = \frac{\exp(\gamma_d + \sum_{p=1}^P \eta_{dp} Y_{ip})}{\sum_{d'=1}^D \exp(\gamma_{d'} + \sum_{p=1}^P \eta_{d'p} Y_{ip})}$$

- $n = 50, P = D = 3,$   
 $\alpha_p \sim \mathcal{N}_{[0, \infty)}(4, 1),$   
 $\beta_p \sim \mathcal{N}_{(0, \infty)}(0.001, 0.1),$   
 $\sigma_p^2 = 5^2, \theta_d^2 = 8^2$
- Two scenarios: “stable increments”,  
 $X = \{50, 100, 150\},$  and  
“increasing increments”,  
 $X = \{50, 100, 300\}$
- $n^* = \{10, 25, 50, 100\}$

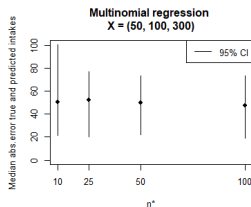
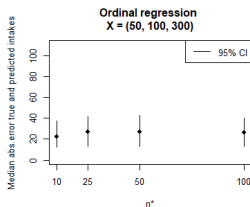
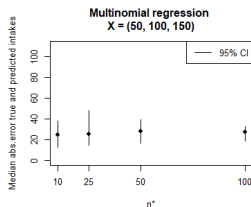
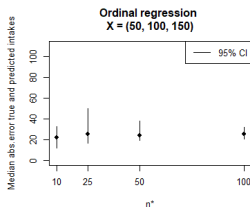




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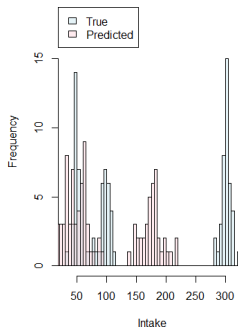


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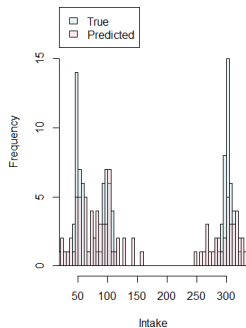
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Multinomial regression  
 $X = (50, 100, 300)$



Ordinal regression  
 $X = (50, 100, 300)$



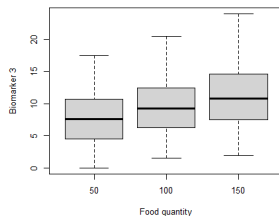
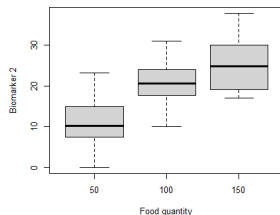
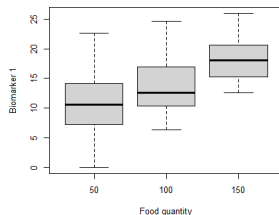
# Simulations - extras

## General settings:

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## Test data:

- **Uniform intakes:**  
 $z_i^* \sim \mathcal{U}(0, 200)$
- **Unbalanced components:** Last component contains most of the observations



# Simulations - extras

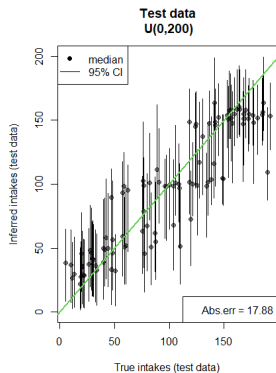
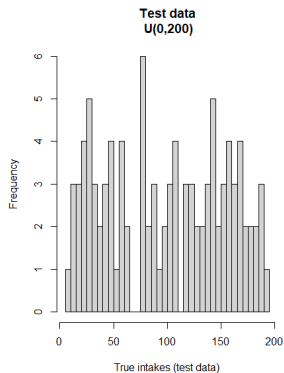
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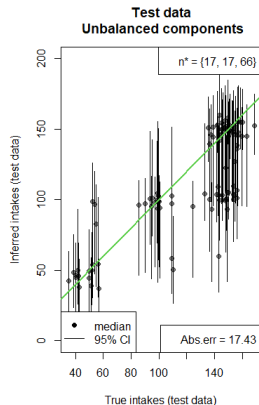
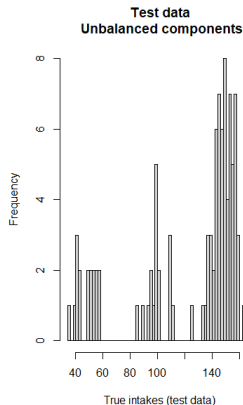
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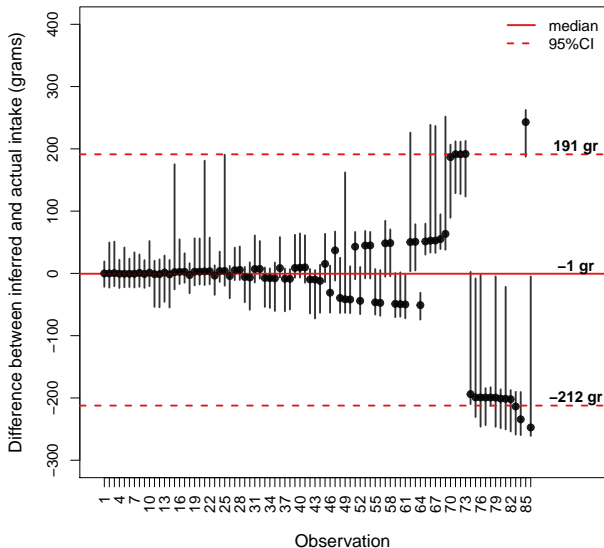


# Apples - predicted intakes

- Leave-one-out  
CV

- **Models:**

- MM
- BLR
- PLS

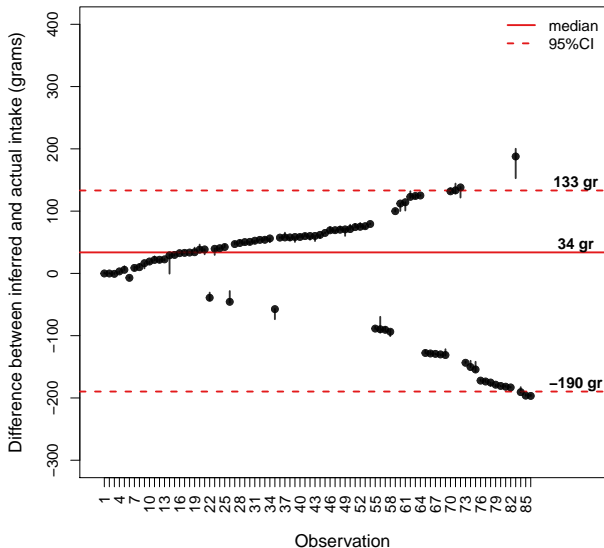


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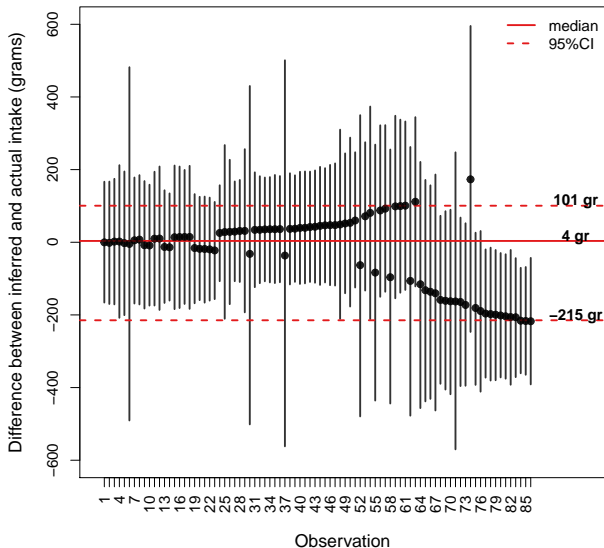


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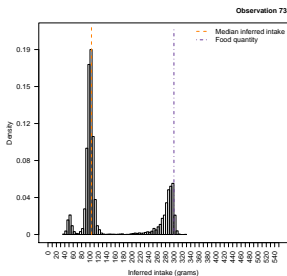
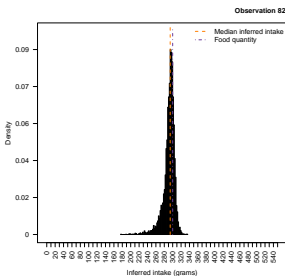
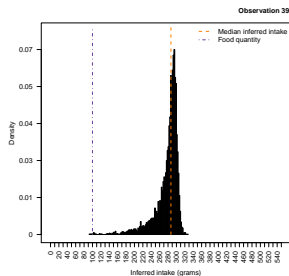
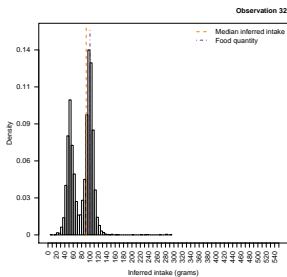
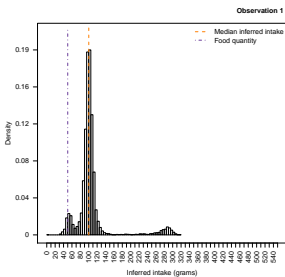
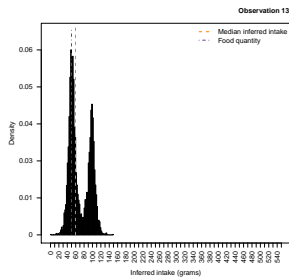
- **Models:**

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- PLS





# Apples - posterior predictive distributions



# Apples - repeated measures?

Data	Dimension	Parameter			
		$\alpha_p$	$\beta_p$	$\sigma_p$	$\theta_d$
Original	1	0.206 (0.267)	0.003 (0.002)	0.353 (0.117)	5.546 (6.514)
	2	0.489 (0.215)	0.005 (0.002)	0.271 (0.104)	7.546 (14.038)
	3	0.612 (0.514)	0.007 (0.004)	0.677 (0.234)	98.989 (71.595)
	4	0.614 (0.323)	0.008 (0.004)	0.382 (0.174)	-
Modified	1	0.214 (0.433)	0.003 (0.003)	0.597 (0.157)	1.806 (1.176)
	2	0.515 (0.416)	0.005 (0.004)	0.600 (0.264)	2.032 (1.773)
	3	0.517 (0.755)	0.009 (0.007)	0.746 (0.258)	8.407 (10.358)
	4	0.662 (0.635)	0.008 (0.006)	0.676 (0.226)	-

## Original data:

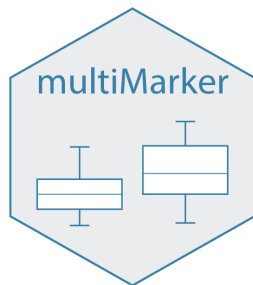
- $n = 86$  observations, treated as independent

## Modified data:

- Each one of the 32 participants appears only once

# Conclusions

- Flexible framework to infer intake from multiple biomarkers
- Uncertainty quantification
- multiMarker: R package and Shiny app
- Easily extendable to other applied contexts, when multiple outcomes are associated with an unobserved variable of interest
- EXTRA: introduction of covariates
- EXTRA: repeated measurements



*Inferring food intake from multiple biomarkers using a latent variable model.* (2021)

D'Angelo, Brennan, Gormley. *Annals of Applied Statistics.*